

Using MCMC and particle filters to forecast stochastic volatility and jumps in financial time series

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Outline of the presentation:

- 1) Motivation
- 2) General asset price process
- 3) MCMC estimation method
- 4) Empirical application of MCMC
- 5) Particle filters
- 6) Calculating forecasts via particle filters

Motivation

- Forecasting volatility and jumps plays a crucial role in many financial applications:
 - Option pricing, VaR calculation, optimal portfolio construction, quantitative trading, etc.
- The main problem is that volatility and jumps are unobservable
- There are currently two classes of models for volatility and jumps estimation and modelling:
 - a) Parametric approach of Stochastic-Volatility Jump-Diffusion models (SVJD) models estimate volatility and jumps as latent state variables with computationally intensive estimation methods such as MCMC.
 - Non-parametric approach using power-variation estimators

 utilize high-frequency data and the asymptotic theory of power variations (Realized Variance, etc.) to construct non-parametric estimates of volatility and jumps

Goal of our research

- Our goal is to develop tools that would enable the estimation and application of more realistic SVJD models for the modelling of asset price behavior
- The areas of study include:
 - Incorporation of additional effects into the SVJD models, such as jump clustering - topic of **Illustration 1**, comes from Fičura and Witzany (2015a), available at: http://papers.ssrn.com/sol3/papers.cfm?abstract_id=2551807
 - Estimation of SVJD models on intraday time series, that would account for the intraday seasonality of volatility and jump intensity – topic of **Illustration 2**, comes from Fičura and Witzany (2015b), available at: http://amse2015.cz/doc/Ficura_Witzany.pdf
 - Utilization of high-frequency power-variation estimators for the estimation of daily SVJD models – Illustration 3, comes from Fičura and Witzany (2015c), available at: https://msed.vse.cz/msed_2015/article/170-Ficura-Milan-paper.pdf



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The general price process

 Let us assume that the logarithmic price of an asset follows a stochastic process defined by the following SDE:

$$dp(t) = \mu(t)dt + \sigma(t)dW(t) + j(t)dq(t)$$

- p(t) is the logarithmic price at time t
- $\mu(t)$ is the instantaneous drift rate
- $\sigma(t)$ is the instantaneous volatility
- W(t) is the Wiener process
- j(t) is a process determining the size of the jumps
- q(t) is a process determining the occurrence of jumps
- $\lambda(t)$ may by a process determining the jump intensity
- We can directly observe only p(t) at discrete points in time
- All of the other processes are unobservable

Stochastic volatility

- Empirical studies show that volatility is time-varying
- i.e. the term $\sigma(t)$ is following its own stochastic process
- A widely used model for $\sigma(t)$ is the log-variance model
- In this model the $h(t) = \ln(\sigma_t^2)$ follows a mean-reverting Ornstein-Uhlenbeck process:

$$dh(t) = \kappa [\theta - h(t)]dt + \xi dW_V(t)$$

- κ determines the strength of the mean reversion
- θ determines the long-term volatility, and
- ξ determines the volatility of volatility
- $W_V(t)$ is a Wiener process that may be correlated with W(t)
- After discretization the O-U process becomes AR(1) process

Self-Exciting jumps

- Empirical studies indicate also some form of jump clustering
- i.e. the intensity of jumps is time varying dq(t)
- We can model the clustering using the self-exciting Hawkes process (with exponential decay function) for the term
- The **jump intensity** $Pr[dq(t) = 1] = \lambda(t)dt$ is then governed by the following process:

$$d\lambda(t) = \kappa[\theta - \lambda(t)]dt + \eta dq(t)$$

• By solving the equation we can get the value of $\lambda(t)$

$$\lambda(t) = \theta + \int_{-\infty}^{t} \eta e^{-\kappa(t-s)} dq(s) = \theta + \sum_{dq(s)=1, s \le t} \eta e^{-\kappa(t-s)}$$

Daily returns and variablity

Assuming the general process for log-price evolution:

$$dp(t) = \mu(t)dt + \sigma(t)dW(t) + j(t)dq(t)$$

• **Daily returns** r(t) = p(t) - p(t-1) can then be expressed:

$$r(t) = \int_{t-1}^{t} \mu(\tau)d\tau + \int_{t-1}^{t} \sigma(\tau)dW(\tau) + \sum_{t-1 \le \tau < t} \kappa(\tau)$$

 The variability of the price process can be expressed with its quadratic variation in the form:

$$QV(t) = \int_{t-1}^{t} \sigma^2(s) ds + \sum_{t-1 \le s < t} \kappa^2(s)$$

• Which is a sum a of integrated variance and jump volatility: QV(t) = IV(t) + IV(t)



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Bayesian estimation methods

- First we have to define the assumed stochastic processes for the logarithmic price, stochastic volatility, jumps, etc.
- Bayesian methods can then be used to estimate the process parameters and the values of the latent state variables
- i.e. the values of the latent stochastic volatilities, jump occurrences, jumps sizes, etc. for every single day in the time series
- Commonly used methods are:
 - Markov Chain Monte Carlo (MCMC)
 - Particle filters (PF)

Markov Chain Monte Carlo

- MCMC algorithm allows us to sample from multivariate distributions by constructing a Markov chain
- It can be used to estimate model parameters and latent state variables by approximating their joint posterior density
- Many versions of the algorithm exist:
 - Gibbs sampler Uses conditional densities to estimate joint density
 - Metropolis-Hastings algorithm Uses rejection sampling
 - Random-walk Uses only the likelihood ratio
 - Multiple-step Similar to Random-Walk but with faster convergence
 - Independence sampling Using approximate densities
- The methods can be combined, using different methods for different variables



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Illustration 1 - The SVJD model

- Model described in Fičura and Witzany (2015a)
- We define our jump-diffusion model with stochastic volatility and self exciting jumps as follows:
- Log-ret: $dp(t) = \mu dt + \sigma(t) dW(t) + j(t) dq(t)$ $j(t) \sim N(\mu_I, \sigma_I)$
- Stoch.vol: $dh(t) = \kappa [\theta h(t)]dt + \xi dW_V(t)$ $h(t) = \ln(\sigma_t^2)$
- Intenzity: $d\lambda(t) = \kappa_J [\theta_J \lambda(t)] dt + \eta_J dq(t)$ $Pr[dq(t) = 1] = \lambda(t) dt$

Euler discretization

- We use the Euler discretization with the assumption that at most 1 jump can happen during one day
- The discrete model has the following equations:

$$r(t) = \mu + \sigma(t)\varepsilon(t) + J(t)Q(t) \qquad V(t) = \sigma_t^2$$

$$h(t) = \alpha + \beta h(t-1) + \gamma \varepsilon_V(t) \qquad h(t) = \ln(V(t))$$

$$\lambda(t) = \alpha_J + \beta_J \lambda(t-1) + \gamma_J Q(t-1) \qquad Q(t) \sim \text{Bern}(\lambda(t))$$

$$\alpha = (1-\beta)\theta \qquad \alpha_J = (1-\beta_J)\theta_J \qquad J(t) \sim N(\mu_J, \sigma_J)$$

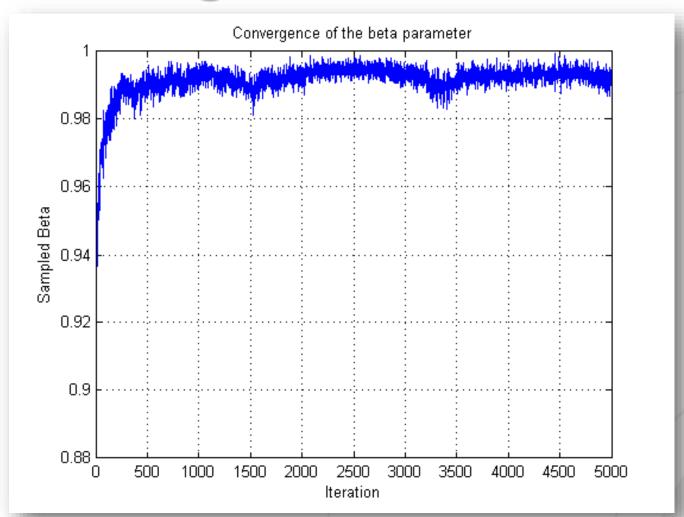
$$\varepsilon(t) \sim N(0,1) \qquad \varepsilon_V(t) \sim N(0,1)$$

- 9 parameters are being estimated: μ , α , β , γ , θ_J , β_J , γ_J , μ_J , σ_J
- And 3 series of latent state variables: V,J,Q

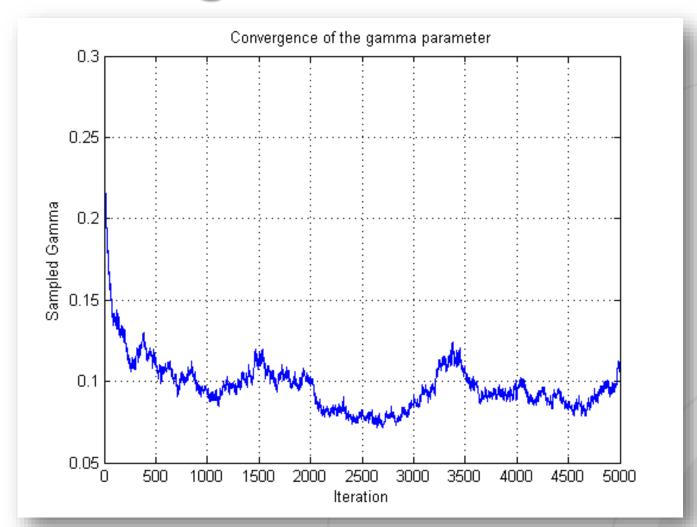
MCMC algorithm

- We estimate the model parameters and the latent state variables using MCMC algorithm
- The MCMC algorithm combines:
 - Gibbs Sampler $(\mu, \sigma, \mu_j, \sigma_j, \alpha, \beta, \gamma, \mathbf{Q}, \mathbf{J})$
 - Accept-Reject Gibbs Sampler (V)
 - Random Walk Metropolis-Hastings $(\theta_I, \beta_I, \gamma_I)$
- The estimation algorithm was firstly tested on simulated time series with mixed results in its ability to identify jumps and their clustering
- Especially if σ_i in the simulated data was not high enough, the jumps mixed with the diffusion volatility and the estimated θ_i was much lower then in the simulations
- We further show only the results for the real data (EUR/USD time series)

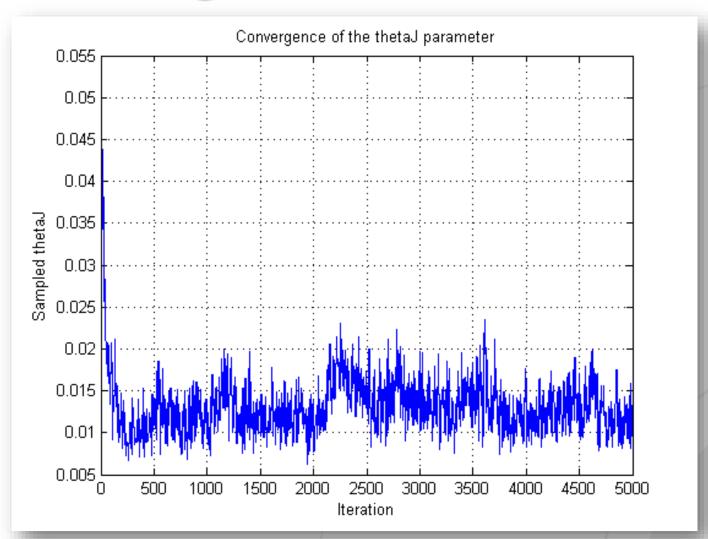
Convergence of Beta



Convergence of Gamma

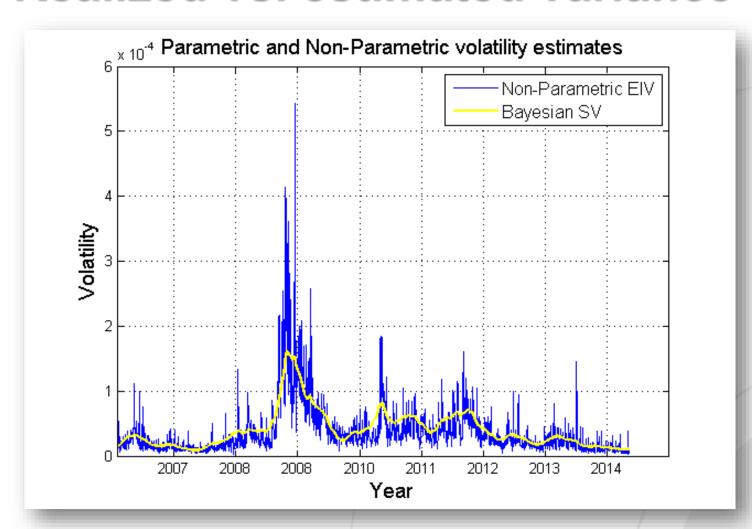


Convergence of ThetaJ

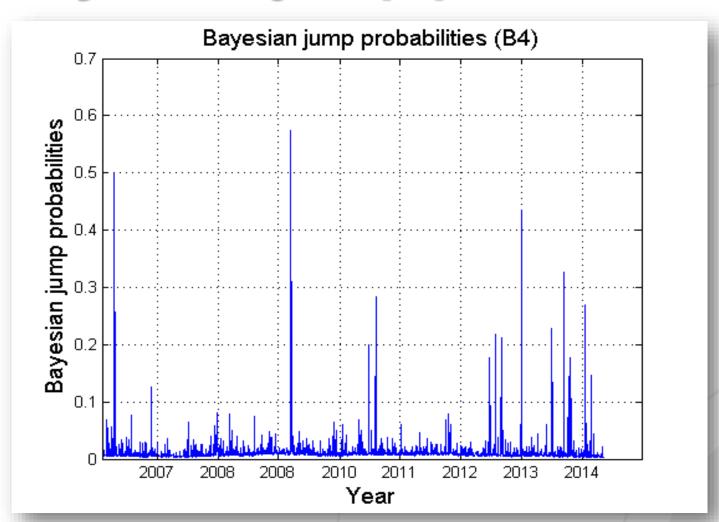




Realized vs. estimated variance

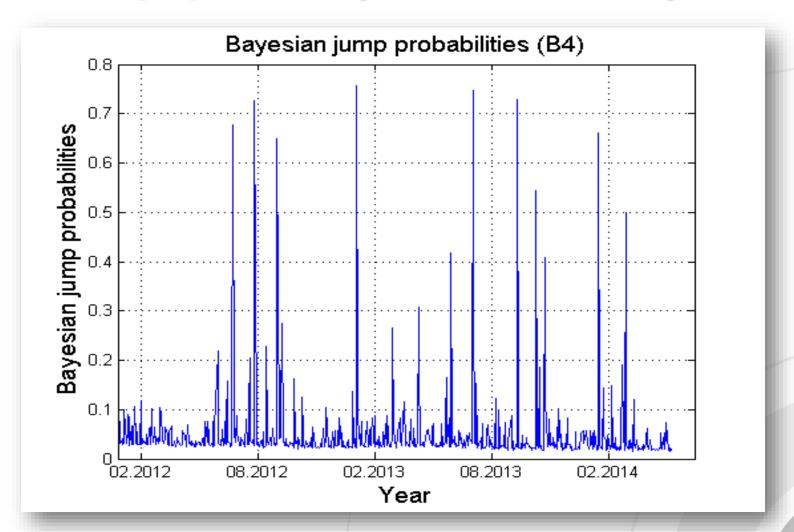


Bayesian jump probabilities





Jump prob. (since 2012)



Can the model identify jumps?

- The performance of the model in this regard does not look very good – too few jumps
- Also the self-exciting property does not seem to be present (the mode of betaJ and gammaJ distributions is close to zero)
- The jumps identified using the bayesian model were further compared with the ones identified using the shrinkage estimator (Z-Statistics):
 - A. Mean probabilities of jumps for every single day were calculated using their bayesian distributions
 - B. The daily **probabilities of jumps** were calculated from the **shrinkage estimator** using the standard normal CDF
- Spearman rank correlation coefficient was calculated between the two variables with value of only <u>0,0171</u>
- i.e. the probabilities of jumps estimated using the two different methods do not seem to be rank correlated

Illustration 2 – Intraday SVJD

- Extended model was used in Fičura and Witzany (2015b), to model volatility and jumps based on intraday price returns (4-hour returns specifically)
- It was necessary to incorporate intraday seasonality of the volatility and jump intensity into the model
- The model contains:
 - 9 parameters related to the stochastic processes
 - 10 parameters associated with the seasonality effects: s_j and λs_{,j} with j going from 1 to 5 (as one in the six seasons is chosen as benchmark with parameter value equal to one)
 - Three vectors of latent state variables: V, J and Q
- The estimation is performed by using a MCMC algorithm combining Gibbs sampler and Metropolis-Hastings algorithm

Full intraday SVJD model:

• The logarithmic returns are given as:

$$r(t) = \mu + \sigma(t)\varepsilon(t) + J(t)Q(t)$$

• The stochastic volatility is given as:

$$\sigma(t) = v(t)s(t)$$

$$h(t) = \alpha + \beta h(t-1) + \gamma \varepsilon_V(t)$$

$$s(t) = \sum_{j=1}^{6} s_j d_j(t)$$

The jump intensity is given as:

$$\lambda(t) = \lambda_H(t)\lambda_S(t)$$

$$\lambda_H(t) = \alpha_J + \beta_J \lambda_H(t-1) + \gamma_J Q(t-1)$$

$$\lambda_{S}(t) = \sum_{j=1}^{6} \lambda_{S,j} d_{j}(t)$$

$$Q(t) \sim Bern[\lambda(t)]$$

$$J(t) \sim N(\mu_J, \sigma_J)$$

$$\varepsilon(t) \sim N(0,1)$$

$$h(t) = \log(V(t))$$

$$V(t) = v^2(t)$$

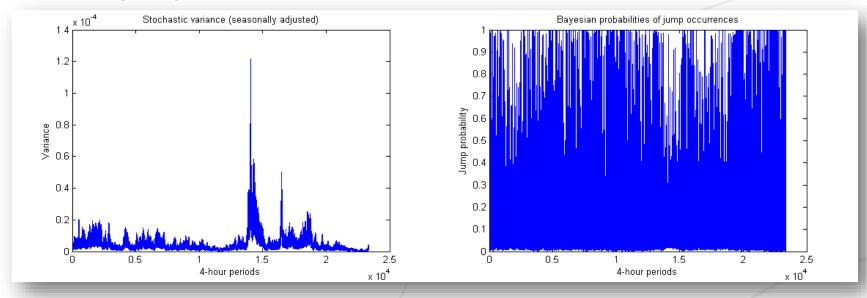
$$\varepsilon_V(t) \sim N(0,1)$$

$$Pr[Q(t)=1]=\lambda(t)$$



Latent state variable estimates

 Posterior mean estimates of the stochastic variances and jump occurrences:

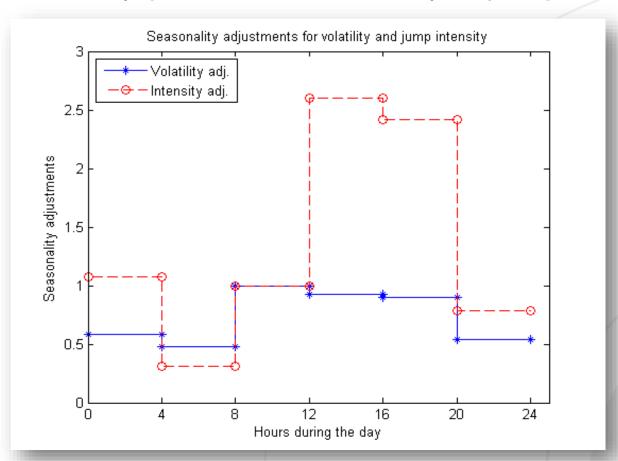


 Range of tests confirmed that the estimates of volatility and jumps based on the intraday SVJD model correspond more closely to the non-parametric estimates of these quantities then when a daily SVJD model is used



Intraday seasonality adjustments

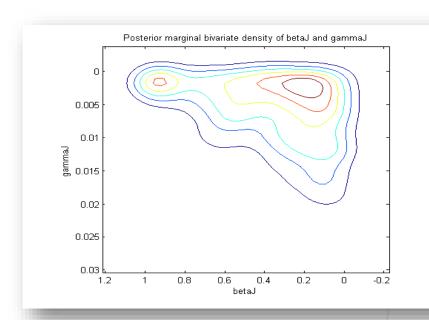
 The model identified the following intraday seasonality patterns of volatility & jump intensity

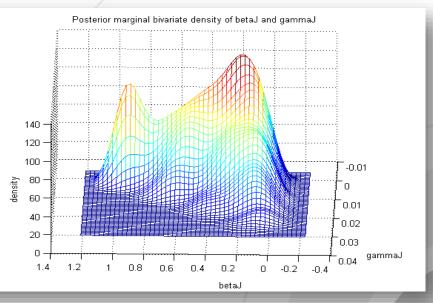




Posterior marginal distribution of betaJ and gammaJ

 The estimated posterior marginal bivariate distribution of betaJ and gammaJ exhibited bimodality







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Particle filters

- Particle filters use weighted set of particles and Bayesian recursion equations in order to estimate the posterior density over a set of latent state-space variables
- Differences compared to MCMC:
- With MCMC we were estimating $p(V_t|\mathcal{F}_T)$
- Where \mathcal{F}_T denotes the information over the whole history of the time series
- With particle filters we are estimating $p(V_t|\mathcal{F}_t)$
- Where \mathcal{F}_t denotes the observable information until time t
- After estimating $p(V_t|\mathcal{F}_t)$, we can make forecasts of $p(V_{t+1}|\mathcal{F}_t)$, $p(V_{t+2}|\mathcal{F}_t)$, etc. via simulations
- The models can thus be used for volatility forecasting, VaR estimation, Option pricing, etc.

Illustration 3 – SVJD-RV-Z

- High-frequency power-variation estimators can be utilized to get additional information for the bayesian estimation of the SVJD models (Fičura and Witzany, 2015c)
- SVJD-RV Model uses realized variance together with daily returns in order to estimate the stochastic volatilities
- SVJD-RV-Z Model uses also the Z-Estimator of jumps in order to more accurately estimate jumps in the time series
- The SVJD-RV-Z model is constructed so that it can distinguish between small jumps (visible only on the intraday frequency) and large jumps (influencing the returns on the daily frequency)
- Models were applied to the EUR/USD exchange rate evolution in the period between 1.11.1999 and 10.10.2014 containing a total of 3 884 trading days

The SVJD-RV-Z model

The logarithmic returns are given as:

$$r(t) = \mu + \sigma(t)\varepsilon(t) + J(t)Q(t)$$

The stochastic volatility is given as:

$$h(t) = \alpha + \beta h(t-1) + \gamma \varepsilon_V(t)$$

The jump intensity is given as:

$$\lambda(t) = \alpha_J + \beta_J \lambda(t-1) + \gamma_J Q(t-1)$$

The realized variance is given as:

$$\log(RV(t) - J^{2}(t)Q(t)) = h(t) + \varepsilon_{RV}(t)$$

• The Z statistics is given as:

$$Z(t) = \mu_Z + \xi_Z Q(t) + \varepsilon_Z$$

$$\varepsilon(t) \sim N(0,1)$$

$$J(t) \sim N(\mu_J, \sigma_J)$$

$$Q(t) \sim Bern[\lambda(t)]$$

$$h(t) = \log(V(t))$$

$$V(t) = \sigma^2(t)$$

$$\varepsilon_V(t) \sim N(0,1)$$

$$Pr[Q(t)=1]=\lambda(t)$$

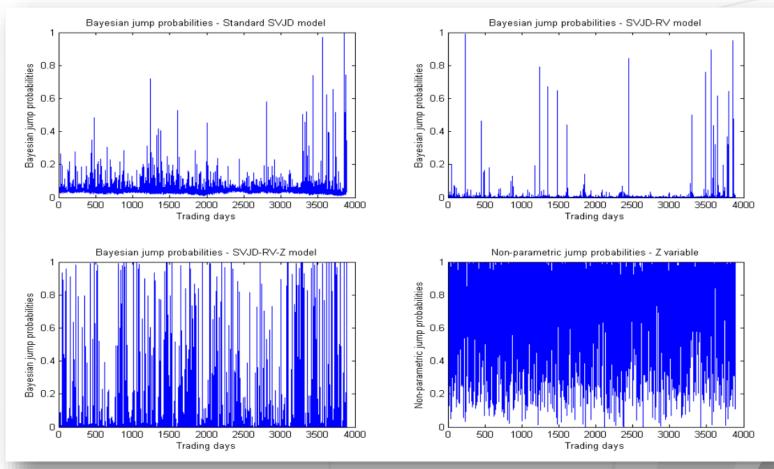
$$\varepsilon_{RV}(t) \sim N(0, \sigma_{RV})$$

$$\varepsilon_z \sim N(0, \sigma_z)$$



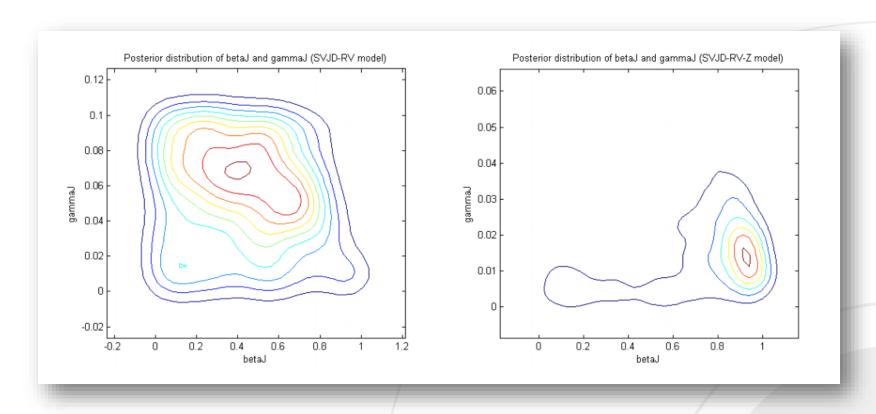
Jump estimates - 4 models

 Estimated jumps under SVJD, SVJD-RV, SVJD-RV-Z (with MCMC) and with the Z-Statistics only





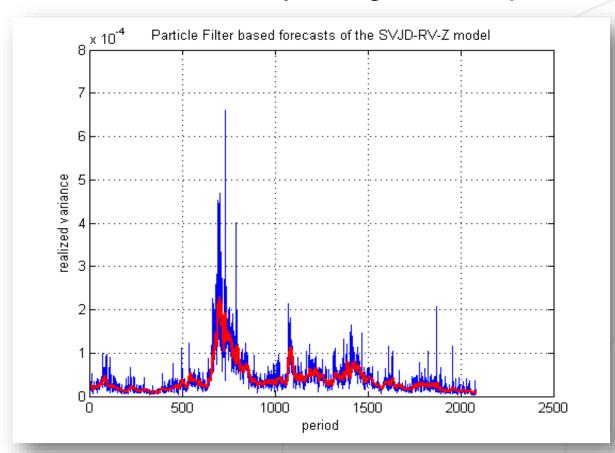
Marginal density of GammaJ and BetaJ (for SVJD-RV and SVJR-RV-Z)





Particle Filter forecasts

 One period ahead forecasts of the SVJD-RV-Z model constructed by using the SIR particle filter



Preliminary results – Out-Sample forecasts

- The out-of-sample R-Squared values are comparable with the HAR model, which is commonly used as benchmark.
- Models estimated with MCMC on the first 1000 periods and applied via particle filters to the rest of the data

Horizon	sv	SV-RV	SVJD	SVJD-RV	SVJD-RV-Z	SVJDH	SVJDH-RV	SVJDH-RV-Z	HAR
1 Day	0.3982	0.4577	0.3867	0.4593	0.4648	0.3923	0.4621	0.4652	0.4746
5 Day	0.5556	0.6523	0.5633	0.6564	0.6539	0.5598	0.6562	0.6628	0.6551
20 Day	0.5825	0.6969	0.5902	0.6986	0.7005	0.5795	0.6955	0.7029	0.7064

Further research:

- We will add jumps in the volatility process
- We will add long-memory to the volatility process
- Auxiliary particle filter instead of SIR filter
- Leverage effect in order to model stock market volatility
- Compare with more advanced modifications of HAR



Conclusions

- The estimation of SVJD models on intraday returns as well as the utilization of intraday power-variation estimators both lead to more realistic jump estimates
- These estimates include significantly more pronounced self-exciting effects of the jumps
- Out-Sample analysis of the models shows that they posses predictive power similar to the HAR model
- As our currently used SVJD models are in principle short-memory, while HAR is a long-memory model, it is expected that the incorporation of long-memory features into the SVJD models may increase the predictive power further
- Similarly the inclusion of volatility jumps and other effects may prove to be useful



Thank you for your attention



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